

# Bayesian Stopping Rules for Meta-analyses

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# What is a meta-analysis?

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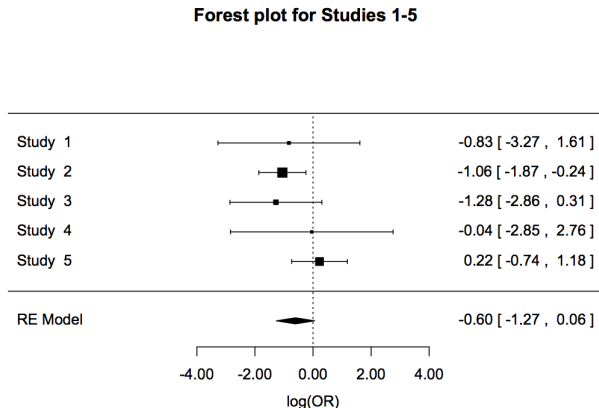
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- Multiple trials are often run on the same question with different populations.
- It would be practical to use this information to determine an overall result for this question.
- A meta-analysis is a statistical tool that allows the analyst to integrate the findings of these trials.
- Analysts use forest plots to show the result of each study and the combined result from the meta-analysis.

# Meta-analysis example

Forest plot for meta-analysis with 5 studies



What happens to the overall estimate of log odds ratio if we add 10 more studies to the meta-analysis?

# Meta-analysis example

Forest plot for meta-analysis with 15 studies

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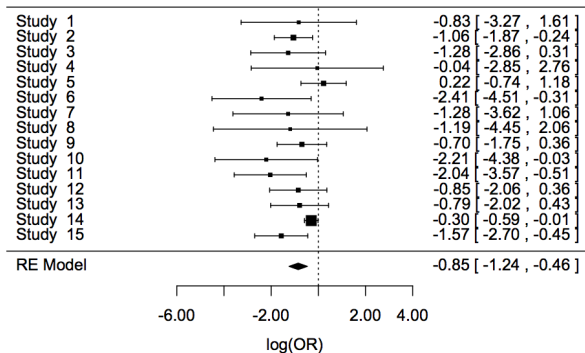
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Forest plot for Studies 1-15



We see that the log odds ratio changed with the addition of 10 more studies. What happens if we add in a mega trial?

# Meta-analysis example

Forest plot for meta-analysis with mega trial

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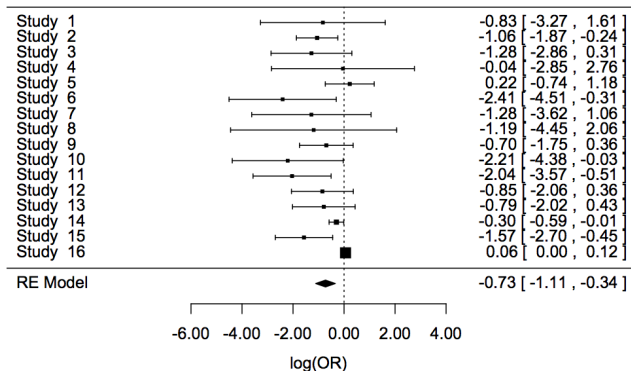
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Forest plot for all 16 studies



Study 16 had a very large impact on the log odds ratio of the meta-analysis. Should we include it in the analysis?

# What is a decision rule?

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- In trials, ethical considerations sometimes require the premature termination of a trial, e.g. treatment being tested may be harmful instead of helpful.
- But stopping the trial early might not give an accurate idea of true treatment effect.
- A stopping rule dictates the earliest moment to terminate the trial without changing the overall conclusion of the trial.

# Bayesian Stopping Rules for trials

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- Stopping rules were originally a frequentist approach.
- Intuitively, it would be reasonable to utilize any existing information concerning the problem at hand (why would we start from nothing if we already know something?).
- Frequentist methods do not incorporate any previous knowledge, but Bayesian methods will.



# Bayesian Statistics

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Let  $\pi(\theta)$  be a probability distribution that incorporates prior information about our parameter of interest  $\theta$ . Let  $L(\theta|x) = \prod_{i=1}^n f(x|\theta)$  be the likelihood of the data. Then the posterior distribution is

$$\pi(\theta|x) = \frac{L(\theta|x)\pi(\theta)}{\int L(\theta|x)\pi(\theta)d\theta} = \frac{L(\theta|x)\pi(\theta)}{m_{\pi}(x)}$$

where  $m_{\pi}(x)$  is the marginal likelihood of the data.

# Greenhouse and Wasserman paper

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- Greenhouse and Wasserman(1995) introduce a robust Bayesian stopping rule for clinical trials.
- Robustness here refers to the use of a class of prior distributions with similar properties instead of being restricted to just one distribution.
- Due to this robustness, their approach is useful for meta-analysis since it requires only a class of priors to be specified, reducing the difficulty in determining an acceptable prior for meta-analysis.

# Greenhouse and Wasserman paper

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- The  $\epsilon$ -contaminated class of priors,  
 $\Gamma_\epsilon = \{\pi = (1 - \epsilon)\pi + \epsilon q : q \in Q\}$  where  $Q$  is some class of prior and  $\epsilon \in (0, 1)$ , allows the analyst to control the contribution of the prior on posterior inference.
- This prior allows us to obtain upper and lower bounds on our inference in the form

$$\bar{E}_{\pi(\theta|x)}(g(\theta)|x) = \sup_{\theta} \frac{(1 - \epsilon)m_\pi(x)E_\pi(g(\theta)|x) + \epsilon L(\theta|x)g(\theta)}{(1 - \epsilon)m_\pi(x) + \epsilon L(\theta|x)}$$

$$\underline{E}_{\pi(\theta|x)}(g(\theta)|x) = \inf_{\theta} \frac{(1 - \epsilon)m_\pi(x)E_\pi(g(\theta)|x) + \epsilon L(\theta|x)g(\theta)}{(1 - \epsilon)m_\pi(x) + \epsilon L(\theta|x)}$$

# Greenhouse and Wasserman paper

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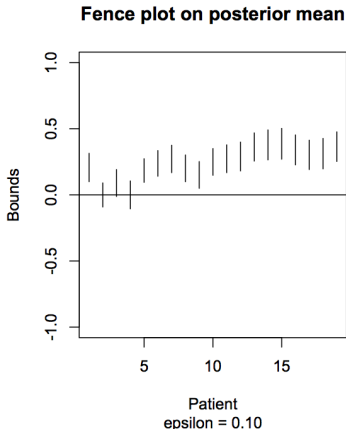
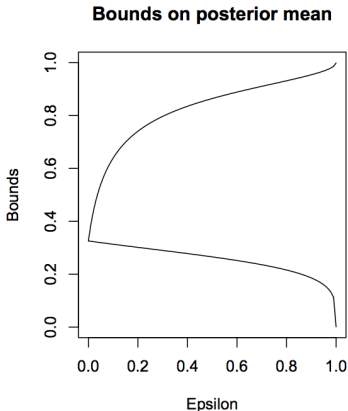
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The form of the posterior bounds also provide nice graphical interpretability.



# Problems

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- The graphical features of the bounds make a wonderful tool for determining stopping rules.
- The largest obstacle in the calculation of the bounds is that the marginal likelihood,  $m_{\pi}(x)$  may not always be in closed form and therefore not analytically integrable.
- Many computational procedures exist for approximating integrals. But how badly will they affect the bounds on our stopping rule?

# BIC approximation

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- The Bayesian Information Criterion (BIC) is usually used for comparing models in model selection problems.
- The standard expression for the BIC is

$$\text{BIC} = -2 \ln L(x|\theta) + k \ln n$$

where  $L(x|\theta)$  is the maximized likelihood,  $k$  is the number of parameters to be estimated, and  $n$  is the sample size.

# BIC approximation

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- There is an alternative form for the BIC that may be used to approximate the marginal likelihood.
- We may write

$$\text{BIC} = -2 \ln L(x|\theta) + k \ln n \approx -2 \ln m_{\pi}(x)$$

where  $m_{\pi}(x)$  is the marginal likelihood

- However, the biggest problem with this approach is that it does not handle random effects well.

# Harvard ECMO clinical trial data

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- Extracorporeal membrane oxygenation (ECMO) compared to conventional medical therapy (CMT) to treat persistent pulmonary hypertension of newborns.
- Patients were randomly assigned to treatment; as soon as 4 deaths occurred in either arm, randomization stopped.
- 19 patients enrolled; 9/9 survived in ECMO arm, 6/10 survived in CMT.



# Harvard ECMO clinical trial data

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- Our likelihood is binomial, so we have conjugate beta priors on the probabilities of success.
- Let  $p_E$  and  $p_C$  be the probability of survival for each arm. We put independent uniform priors on each which can be expressed in terms of a Beta distribution
- Here, we can obtain the marginal likelihood in closed form.
- We are in a fixed effect setting, so we would expect our approximations to be comparable, depending on the starting values of the function.

# Harvard ECMO clinical trial data

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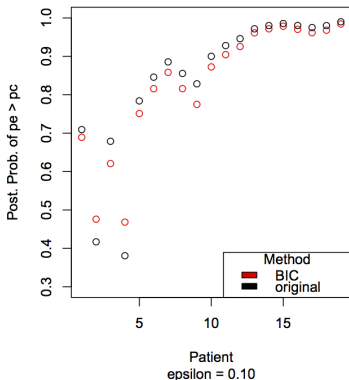
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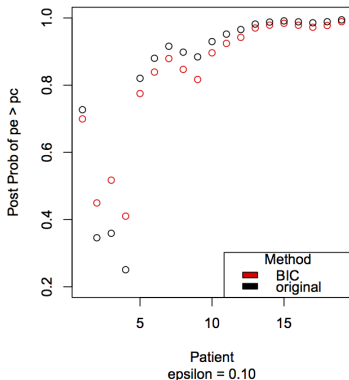
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Here we see the difference in posterior probabilities under both methods.

Contaminated posterior under post. mean

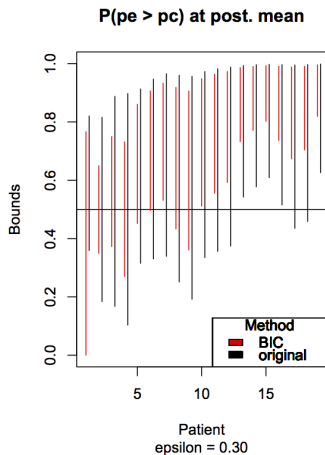
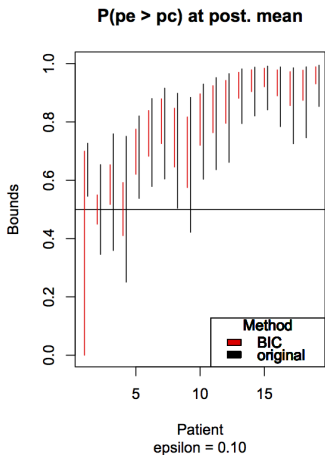


Contaminated posterior under post. mode



# Harvard ECMO clinical trial data

Now we can see the effect each method would have on the stopping decision in the trial.



# Harvard ECMO clinical trial data

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- The BIC approximation of the marginal likelihood for trial data seems to underestimate the bounds of the posterior probabilities.
- If we are quite confident in our choice of prior ( $\epsilon$  small), then we seem to agree on a stopping point with the closed form method.
- However, as we become more uncertain of the validity of the prior, the underestimation of the BIC becomes problematic.
- Small sample size could be to blame; so what happens if we apply this technique to a meta-analysis?

# Meta-analysis simulation study

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- Simulate 16 studies of size 200 each, with 100 subjects in each of the treatment and control arms.
- The number of cases in each arm were generated from a Binomial distribution where

$$\log(p_p) \sim U(-3.665, -0.995)$$

$$\log(p_t) = \log(p_p) + \theta$$

where  $\theta$  is the treatment effect.

- This allows us to have a common baseline among the studies but a study-specific treatment effect.

# Meta-analysis simulation study

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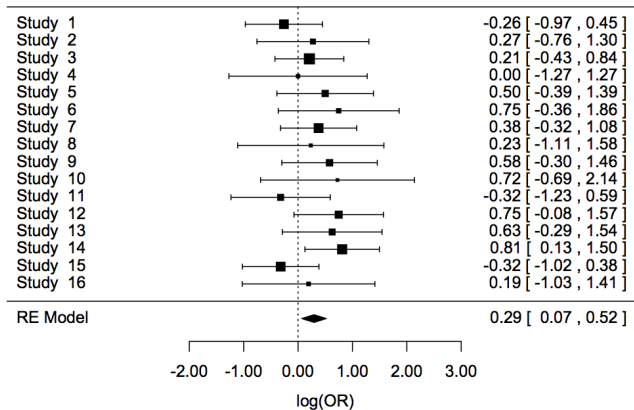
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Forest plot for the simulated data.



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- While allowing for a random effect is far more realistic, we can no longer employ our previous methods to these data.
- The marginal likelihood can no longer be obtained in closed form.
- The BIC approximation is for fixed effect models only where it does a questionable job already.
- So what can we do?

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- Linear mixed models (LM) can incorporate a random effect in the model and provide us with an estimate of the BIC.
- Random effects models are far more realistic as they allow modelling of both between- study and within-study variability.
- Fitting the simulated data with "lmer" in R, we can extract an estimate of the marginal likelihood under a random effects model and a value for the log likelihood, which we can use in our calculations of the posterior probabilities.



# Meta-analysis simulation study

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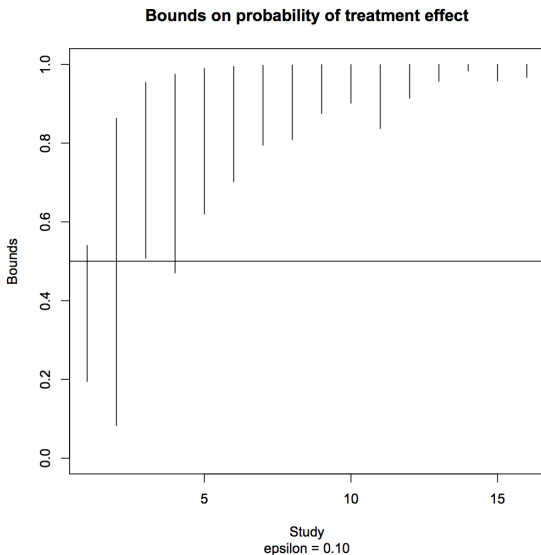
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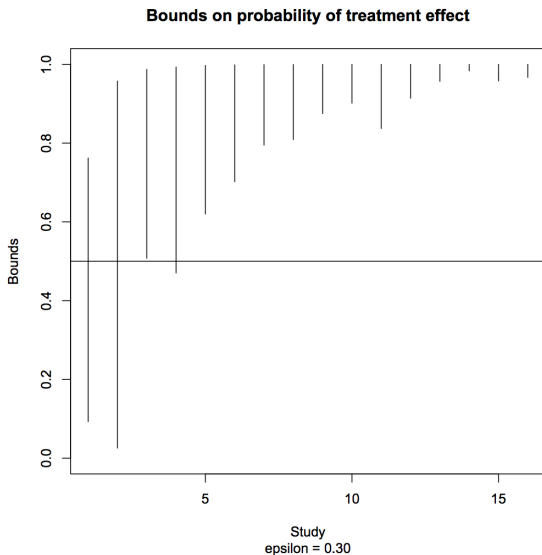
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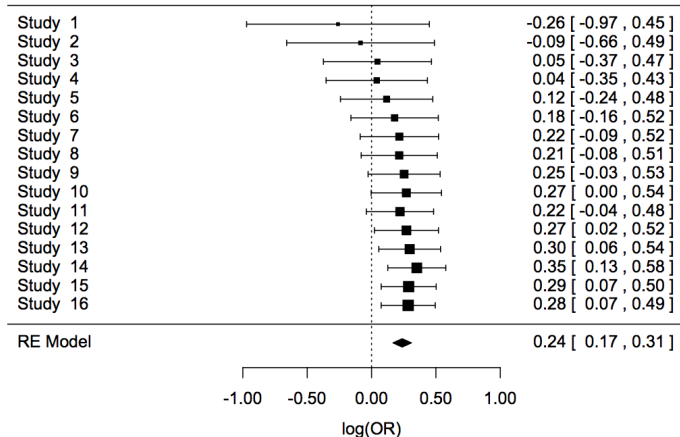
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## Forest plot for cumulative meta-analysis



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- Through the bounds that we obtain on our posterior probability of treatment effect, we see that even under the worst choice of prior, we would need a very high cutoff for probability of treatment effect in order to conclude that all 16 studies were necessary.
- We notice that as we increase our epsilon, we do not notice much of a change in the bounds.
- Likely caused by the size of each study: at 5 studies, we have a cumulative sample size of 500 subjects in each arm.

# Conclusion

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- For fixed effect models, we can use the naive BIC approximation for our marginal likelihood and it performs adequately.
- For random effect models, the problem becomes more complicated, but it appears that the BIC value from fitting a linear mixed model can be used.
- The  $\epsilon$ -contaminated approach provides a nice technique for managing the effect of our prior distribution on the posterior inference.
- Added advantage: nice graphical methods for determining stopping rules.

# Future work

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- Use path sampling method to approximate the marginal likelihood.
- Path sampling is a Monte Carlo computation approach, frequently use in Bayesian statistics as an integration technique used to estimate marginal likelihoods.
- Also investigate better asymptotic approximations based on ML estimates as discussed in Pauler et al. (1999).